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Survival analysis of a stochastic cooperation system with functional response in a polluted environment

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Abstract

In this paper, we propose and study a stochastic two-species cooperation model with functional response in a polluted environment. We first perform the survival analysis and establish sufficient conditions for extinction, weak persistence, and stochastic permanence. Then we further perform the survival analysis based on the temporal average of population size and derive sufficient conditions for the strong persistence in the mean and weak persistence in the mean. Finally, we present numerical simulations to justify the theoretical results.

Keywords: Cooperation system; White noise; Polluted environment; Survival analysis; Numerical simulations

1 Introduction

With rapid development of industries and agriculture, a mass of toxicants has been emitted into the environment, such as the industrial wastewater, domestic sewage, and other contaminants. The presence of a variety of toxicants in the environment not only seriously threatened the survival of the exposed populations but also affected the human life style. Therefore it is important to estimate the environmental toxicity, which requires quantitative estimates for the survival risk of species in a polluted environment. This motivates scholars to utilize the mathematical models to assess the effects of toxicants on various ecosystems. Hallam et al. [1–3] did pioneering works. It has been a significant topic of considerable researches, and more and more deterministic models have been proposed and analyzed (see [4–17]). They all provide a great insight into the effects of pollutants. In this paper, we mainly attempt to study a two-species cooperation model with functional response. We assume that the environment is of complete spatial homogeneity and there is no migration. Let $x_i(t)$ represent the population size (population density) of the *i*th species at time *t*. We consider the following cooperation model in a polluted environment:

$$\begin{cases} x_1'(t) = x_1(t)[r_1 - a_1x_1(t) + \frac{c_2x_2(t)}{1 + b_2x_2(t)}], \\ x_2'(t) = x_2(t)[r_2 - a_2x_2(t) + \frac{c_1x_1(t)}{1 + b_1x_1(t)}], \end{cases}$$
(1.1)

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where the positive coefficients r_1, r_2 and a_1, a_2 are the intrinsic growth rates and selfinhibition rates, respectively. In the classical cooperation model the mutualism effects are described by a bilinear function, that is, x_i response to x_j is assumed to be increasingly monotonic, an inherent assumption meaning that the more x_j there exist in the environment, the better off the x_i . But here the term $c_i x_i(t)/(1 + b_i x_i(t))$ represents the functional response function, and moreover, it is an increasing function with respect to x_i and has a saturation value for large enough x_i . The positive coefficients c_j measure the mutualism effects of species x_j on species x_i , and b_j , i, j = 1, 2, $i \neq j$, are positive control constants. For a relevant ecological model of system (1.1), we refer the readers to [18].

Now we are in the position to describe the dynamics of population in a polluted environment. We assume that the living organisms absorb part of toxicants into their bodies, the dynamics of the population is affected by internal toxicant, and the individuals in two species have identical concentration of organismal toxicant at time t (see [5]). To simplify the mathematical model, we assume that the capacity of the environment is so large that the change of toxicant in the environment that comes from uptake and egestion by the organisms can be neglected (see [3]). Let $C_0(t)$ represent the concentration of toxicant in the environment at time t, and let $C_E(t)$ represent the concentration of toxicant in the environment at time t. A coupling between species and toxicant is formulated by assuming that the intrinsic growth rate of the ith population, $r_{i0} - r_{i1}C_0(t)$, is a linear function of concentration of toxicant present in the organism. Then the population dynamics can be given as follows:

$$\begin{cases} x_1'(t) = x_1(t)[r_{10} - r_{11}C_0(t) - a_1x_1(t) + \frac{c_2x_2(t)}{1 + b_2x_2(t)}], \\ x_2'(t) = x_2(t)[r_{20} - r_{21}C_0(t) - a_2x_2(t) + \frac{c_1x_1(t)}{1 + b_1x_1(t)}], \\ \frac{dC_0(t)}{dt} = dC_E(t) + \eta - (l_1 + l_2)C_0(t), \\ \frac{dC_E(t)}{dt} = -hC_E(t) + u(t), \end{cases}$$
(1.2)

where r_{i0} is the intrinsic growth rate of the *i*th species in the absence of toxicant, r_{i1} , i = 1, 2, is the dose-response rate of species *i* to the toxicant concentration, *d* represents the uptake rate of toxicant from environment per unit biomass, η represents the intake rate of toxicant from food per unit biomass, l_1 represents the organismal net ingestion rates of toxicant, l_2 represents the organismal excretion rates of toxicant, and *h* represents the loss rate of toxicant from the environment due to the processes such as biological transformation, chemical hydrolysis, volatilization, microbial degradation, and photosynthetic degradation. Here u(t) represents the input rate of exogenous toxin at time *t*, which is a nonnegative continuous function defined on $[0, \infty)$ with $u_1 := \sup_{t>0} u(t) > 0$.

Considering that the fate of young recruits after reproduction is quite sensitive to the environmental fluctuations, so the growth rate of population is inevitably affected. May and Allen [19] have claimed that the growth rates in population systems should be fluctuating around some average values because of the environmental fluctuations. In this sense, it is necessary to incorporate the environmental fluctuations into our model. In practice, we usually estimate the growth rate r_{i0} by an average value plus an error term; by the central limit theorem the error term follows a normal distribution, and thus we can approximate the error term by $\sigma_i \dot{B}_i(t)$, that is,

$$r_{i0} \rightarrow r_{i0} + \sigma_i \dot{B}_i(t), \quad i = 1, 2,$$

where $\dot{B}_i(t)$ represents a white noise process (i.e., $B_i(t)$ is a standard Brownian motion), and σ_i^2 is the intensity of the white noise. According to our discussion, a stochastic two-species cooperation model in polluted environment is derived as follows:

$$\begin{cases} dx_{1}(t) = x_{1}(t)[r_{10} - r_{11}C_{0}(t) - a_{1}x_{1}(t) + \frac{c_{2}x_{2}(t)}{1 + b_{2}x_{2}(t)}] dt + \sigma_{1}x_{1}(t) dB_{1}(t), \\ dx_{2}(t) = x_{2}(t)[r_{20} - r_{21}C_{0}(t) - a_{2}x_{2}(t) + \frac{c_{1}x_{1}(t)}{1 + b_{1}x_{1}(t)}] dt + \sigma_{2}x_{2}(t) dB_{2}(t), \\ \frac{dC_{0}(t)}{dt} = dC_{E}(t) + \eta - (l_{1} + l_{2})C_{0}(t), \\ \frac{dC_{E}(t)}{dt} = -hC_{E}(t) + u(t). \end{cases}$$
(1.3)

The initial conditions satisfy the conditions

$$x_1(0) > 0,$$
 $x_2(0) > 0,$ $0 \le C_0(0) < 1,$ $0 \le C_E(0) < 1.$

In recent years, many interesting and important works about the stochastic model in polluted environment have been reported (see [20-33]). But to the best of our knowledge, there exist few published papers concerning system (1.3). Noting that the last two equations in system (1.3) can be explicitly solved, so we only need to consider the following subsystem:

$$\begin{cases} dx_1(t) = x_1(t)[r_{10} - r_{11}C_0(t) - a_1x_1(t) + \frac{c_2x_2(t)}{1 + b_2x_2(t)}] dt + \sigma_1x_1(t) dB_1(t), \\ dx_2(t) = x_2(t)[r_{20} - r_{21}C_0(t) - a_2x_2(t) + \frac{c_1x_1(t)}{1 + b_1x_1(t)}] dt + \sigma_2x_2(t) dB_2(t). \end{cases}$$
(1.4)

Motivated by the existing results, the rest of this paper is arranged as follows. In Sect. 2, we introduce several commonly used basic lemmas and classical definitions. In Sect. 3, we perform the survival analysis and establish sufficient criteria for extinction, weak persistence, and stochastic permanence. In Sect. 4, we further discuss the survival analysis based on the temporal average of population size and derive sufficient conditions for the strong persistence in the mean and weak persistence in the mean. In Sect. 5, we present several numerical simulations to validate our theoretical results. The limitation of the model is also discussed in the last section.

2 Preliminaries

From now on, unless otherwise specified, we always work on the complete probability space (Ω, \mathcal{F}, P) with filtration $\{\mathcal{F}_t\}_{t\geq 0}$ satisfying the usual conditions, that is, it is right continuous, and \mathcal{F}_0 contains all *P*-null sets. Note that both $C_0(t)$ and $C_E(t)$ represent the concentrations of toxicant, so to be realistic, we must have $0 \leq C_0(t) < 1$ and $0 \leq C_E(t) < 1$ for all $t \geq 0$. In fact, we can prove that by solving the last two equations of system (1.3).

For convenience and simplicity, we introduce the following notations:

$$\langle f(t) \rangle \triangleq t^{-1} \int_0^t f(s) \, ds, \qquad \langle f \rangle_* \triangleq \liminf_{t \to \infty} t^{-1} \int_0^t f(s) \, ds,$$

$$\langle f \rangle^* \triangleq \limsup_{t \to \infty} t^{-1} \int_0^t f(s) \, ds, \qquad f_* \triangleq \liminf_{t \to \infty} f(t), \qquad f^* \triangleq \limsup_{t \to \infty} f(t).$$

Definition 2.1 (See [34]) To state and prove our main results, we recall some classical concepts.

- Population x_i is said to be extinct if $\lim_{t\to\infty} x_i(t) = 0$, i = 1, 2, a.s.
- Population x_i is said to be weakly persistent if $\limsup_{t\to\infty} x_i(t) > 0$ a.s.
- Population x_i is said to be weakly persistent in the mean if $\limsup_{t\to\infty} \langle x_i(t) \rangle > 0$ a.s.
- Population x_i is said to be strongly persistent in the mean if $\liminf_{t\to\infty} \langle x_i(t) \rangle > 0$ a.s.
- System (1.4) is said to be stochastically permanent if for every $\varepsilon \in (0, 1)$, there exists a pair of positive constants $\alpha = \alpha(\varepsilon)$, $\beta = \beta(\varepsilon)$ such that for any positive initial condition x(0), the solution x(t) satisfies

$$\liminf_{t \to \infty} P\{x_i(t) \ge \alpha\} \ge 1 - \varepsilon, \qquad \liminf_{t \to \infty} P\{x_i(t) \le \beta\} \ge 1 - \varepsilon.$$
(2.1)

3 Survival analysis

In this section, we first study the existence and uniqueness of a globally positive solution for the biological significance and then perform the survival analysis of system (1.4).

Theorem 3.1 For every initial value $x(0) = (x_1(0), x_2(0)) \in R^2_+$, system (1.4) has a unique solution x(t), and x(t) remains in R^2_+ for all $t \ge 0$ with probability one.

Proof The proof of the theorem is standard. Firstly, let us consider the following stochastic system:

$$\begin{cases} dy_1(t) = [r_{10} - 0.5\sigma_1^2 - r_{11}C_0(t) - a_1e^{y_1(t)} - \frac{c_2e^{y_2(t)}}{1 + b_2e^{y_2(t)}}]dt + \sigma_1 dB_1(t), \\ dy_2(t) = [r_{20} - 0.5\sigma_2^2 - r_{21}C_0(t) - a_2e^{y_2(t)} - \frac{c_1e^{y_1(t)}}{1 + b_1e^{y_1(t)}}]dt + \sigma_2 dB_2(t), \end{cases}$$
(3.1)

with initial condition $y_1(0) = \ln x_1(0)$, $y_2(0) = \ln x_2(0)$. It is easy to verify that the coefficients of system (3.1) satisfy the local Lipschitz condition, and thus system (3.1) has a unique solution y(t) on $[0, \tau_e)$, where τ_e is the explosion time (see [35]). By applying Itô's formula and a simple calculation, it is easy to derive that $x_1(t) = e^{y_1(t)}$, $x_2(t) = e^{y_2(t)}$ is the unique positive local solution to system (1.4) with initial value $x(0) \in R^2_+$. To show that this solution is global, we only need to show that $\tau_e = \infty$.

Let n_0 be sufficiently large such that every component of x(0) remains in the interval $[\frac{1}{m_0}, n_0]$. For each integer $n \ge n_0$, we define the stopping time

$$\tau_n = \inf \left\{ t \in [0, \tau_e) : x_1(t) \notin \left(\frac{1}{n}, n\right) \text{ or } x_2(t) \notin \left(\frac{1}{n}, n\right) \right\}.$$

Here we set $\inf \emptyset = +\infty$ (\emptyset denotes the empty set). Obviously, τ_n is increasing as $n \to \infty$. Let $\tau_{\infty} = \lim_{n \to \infty} \tau_n$, whence $\tau_{\infty} \le \tau_e$ almost surely; if we can show that $\tau_{\infty} = \infty$ almost surely, then $\tau_e = \infty$ almost surely, and the proof is completed.

Define the C^2 -function $\widetilde{V}(x) : \mathbb{R}^2_+ \to \mathbb{R}_+$ as

$$\widetilde{V}(x_1, x_2) = x_1 - 1 - \ln x_1 + x_2 - 1 - \ln x_2.$$

The nonnegativity of this function can be seen from $v - 1 - \ln v \ge 0$ for v > 0. Let T > 0 be an arbitrary constant. Then for $0 \le t \le \tau_n \land T$, by Itô's formula we obtain

$$d\widetilde{V}(x_1, x_2)$$

$$= \left(1 - \frac{1}{x_1}\right) dx_1 + 0.5 \frac{1}{x_1^2} (dx_1)^2 + \left(1 - \frac{1}{x_2}\right) dx_1 + 0.5 \frac{1}{x_2^2} (dx_2)^2$$

$$= (x_1 - 1) \left[r_{10} - r_{11}C_0(t) - a_1x_1 + \frac{c_2x_2}{1 + b_2x_2} \right] dt + (x_1 - 1)\sigma_1 dB_1(t) + \frac{1}{2}\sigma_1^2 dt$$

$$+ (x_2 - 1) \left[r_{20} - r_{21}C_0(t) - a_2x_2 + \frac{c_1x_1}{1 + b_1x_1} \right] dt + (x_2 - 1)\sigma_2 dB_2(t) + \frac{1}{2}\sigma_2^2 dt$$

$$= \left\{ -a_1x_1^2 + (r_{10} + a_1)x_1 - r_{11}C_0(t)x_1 + \frac{c_2x_1x_2}{1 + b_2x_2} - \frac{c_2x_2}{1 + b_2x_2} + \frac{1}{2}\sigma_1^2 - r_{10} + r_{11}C_0(t) - a_2x_2^2 + (r_{20} + a_2)x_2 - r_{21}C_0(t)x_2 + \frac{c_1x_1x_2}{1 + b_1x_1} - \frac{c_1x_1}{1 + b_1x_1} + \frac{1}{2}\sigma_2^2 - r_{20} + r_{21}C_0(t) \right\} dt + (x_1 - 1)\sigma_1 dB_1(t) + (x_2 - 1)\sigma_2 dB_2(t)$$

$$\leq \left\{ -a_1x_1^2 + \left(r_{10} + a_1 + \frac{c_2}{b_2}\right)x_1 + \frac{1}{2}\sigma_1^2 + r_{11} - r_{10} - a_2x_2^2 + \left(r_{20} + a_2 + \frac{c_1}{b_1}\right)x_2 + \frac{1}{2}\sigma_2^2 + r_{21} - r_{20} \right\} dt + (x_1 - 1)\sigma_1 dB_1(t) + (x_2 - 1)\sigma_2 dB_2(t)$$

$$= G(x_1, x_2) dt + (x_1 - 1)\sigma_1 dB_1(t) + (x_2 - 1)\sigma_2 dB_2(t), \qquad (3.2)$$

where

$$G(x_1, x_2) = -a_1 x_1^2 + \left(r_{10} + a_1 + \frac{c_2}{b_2}\right) x_1 + \frac{1}{2} \sigma_1^2 + r_{11} - r_{10}$$
$$-a_2 x_2^2 + \left(r_{20} + a_2 + \frac{c_1}{b_1}\right) x_2 + \frac{1}{2} \sigma_2^2 + r_{21} - r_{20}.$$
(3.3)

Clearly, $G(x_1, x_2)$ is upper bounded, say by \mathcal{M} . So we have

$$d\widetilde{V}(x_1, x_2) \le \mathcal{M} \, dt + (x_1 - 1)\sigma_1 \, dB_1(t) + (x_2 - 1)\sigma_2 \, dB_2(t).$$
(3.4)

Integrating both sides from 0 to $\tau_n \wedge T$, we have

$$\int_{0}^{\tau_{n}\wedge T} d\widetilde{V}(x_{1}, x_{2}) \leq \int_{0}^{\tau_{n}\wedge T} \mathcal{M} dt + \int_{0}^{\tau_{n}\wedge T} (x_{1} - 1)\sigma_{1} dB_{1}(t) + \int_{0}^{\tau_{n}\wedge T} (x_{2} - 1)\sigma_{2} dB_{2}(t).$$
(3.5)

Taking the expectations on both sides yields that

$$E\left[\widetilde{V}(x_1(\tau_n \wedge T), x_2(\tau_n \wedge T))\right] \leq \widetilde{V}(x_1(0), x_2(0)) + \mathcal{M}E(\tau_n \wedge T)$$
$$\leq \widetilde{V}(x_1(0), x_2(0)) + \mathcal{M}T.$$
(3.6)

Let $\Omega_n = \{\tau_n \leq T\}$. For arbitrary $\omega \in \Omega_n$, there exists some *i* such that $x_i(\tau_n, \omega)$ equals either *n* or $\frac{1}{n}$, and thus $\widetilde{V}(x(\tau_n, \omega))$ is not less than $(n - 1 - \ln n) \wedge (\frac{1}{n} - 1 + \ln n) \triangleq \mu(n)$. It then follows from (3.6) that

$$\mu(n)P(\tau_n \leq T) \leq E\left[1_{\Omega_n}(\omega)\widetilde{V}(x_1(\tau_n,\omega),x_2(\tau_n,\omega))\right]$$

$$\leq \widetilde{V}(x_1(0), x_2(0)) + \mathcal{M}T, \tag{3.7}$$

where 1_{Ω_n} is the indicator function of Ω_n . Let $n \to \infty$. Then $\mu(n) \to \infty$, and (3.7) leads to $P(\tau_{\infty} \leq T) = 0$. By the arbitrariness of T we have $P(\tau_{\infty} = \infty) = 1$ almost surely. This completes the proof of Theorem 3.1.

It is well known that the threshold is very important for assessing the risk of extinction for species exposed to toxicant from a biological point of view. In the following, we first show that the *p*th moment of the solution of system (1.4) is upper bounded and then establish the threshold between weak persistence and extinction for species x_i modeled by (1.4). To begin with, we present the fundamental assumption that $r_{i0} > 0.5\sigma_i^2$. Unless otherwise stated, it is always assumed in this paper.

Theorem 3.2 For any p > 1, there exists a positive constant K = K(p) such that the solution x(t) of system (1.4) has the property that

$$\limsup_{t \to \infty} E |x(t)|^p \le K.$$
(3.8)

Proof Define $V(x_1, x_2) = x_1^p + x_2^p$. By Itô's formula, we have

$$\begin{aligned} dV(x_{1}, x_{2}) \\ &= px_{1}^{p-1} dx_{1} + 0.5p(p-1)x_{1}^{p-2}(dx_{1})^{2} + px_{2}^{p-1} dx_{2} + 0.5p(p-1)x_{2}^{p-2}(dx_{2})^{2} \\ &= px_{1}^{p} \left(r_{10} - r_{11}C_{0}(t) - a_{1}x_{1} + \frac{c_{2}x_{2}}{1 + b_{2}x_{2}} \right) dt + p\sigma_{1}x_{1}^{p} dB_{1}(t) + 0.5p(p-1)\sigma_{1}^{2}x_{1}^{p} dt \\ &+ px_{2}^{p} \left(r_{20} - r_{21}C_{0}(t) - a_{2}x_{2} + \frac{c_{1}x_{1}}{1 + b_{1}x_{1}} \right) dt + p\sigma_{2}x_{2}^{p} dB_{2}(t) + 0.5p(p-1)\sigma_{2}^{2}x_{2}^{p} dt \\ &= \left\{ x_{1}^{p} \left(1 + r_{10}p - r_{11}pC_{0}(t) - a_{1}px_{1} + p\frac{c_{2}x_{2}}{1 + b_{2}x_{2}} + 0.5p(p-1)\sigma_{1}^{2} \right) + x_{2}^{p} \left(1 + r_{20}p - r_{21}pC_{0}(t) - a_{2}px_{2} + p\frac{c_{1}x_{1}}{1 + b_{1}x_{1}} + 0.5p(p-1)\sigma_{2}^{2} \right) - V(x_{1},x_{2}) \right\} dt \\ &+ p\sigma_{1}x_{1}^{p} dB_{1}(t) + p\sigma_{2}x_{2}^{p} dB_{2}(t) \\ &\leq \left\{ x_{1}^{p} \left(1 + r_{10}p + \frac{c_{2}p}{b_{2}} + 0.5p(p-1)\sigma_{1}^{2} - a_{1}px_{1} \right) + x_{2}^{p} \left(1 + r_{20}p + \frac{c_{1}p}{b_{1}} \right) \right\} \\ &+ 0.5p(p-1)\sigma_{2}^{2} - a_{2}px_{2} \right) - V(x_{1},x_{2}) \right\} dt + p\sigma_{1}x_{1}^{p} dB_{1}(t) + p\sigma_{2}x_{2}^{p} dB_{2}(t) \end{aligned}$$

where

$$L(x_1, x_2) = x_1^p \left(1 + r_{10}p + \frac{c_2p}{b_2} + 0.5p(p-1)\sigma_1^2 - a_1px_1 \right) + x_2^p \left(1 + r_{20}p + \frac{c_1p}{b_1} + 0.5p(p-1)\sigma_2^2 - a_2px_2 \right).$$
(3.10)

Obviously, $L(x_1, x_2)$ is upper bounded; we denote it by \mathcal{L} , that is,

$$\mathcal{L} \triangleq \sum_{i=1}^{2} \frac{1 + r_{i0}p + \frac{c_{i}p}{b_{j}} + 0.5p(p-1)\sigma_{i}^{2}}{(p+1)a_{i}} > 0.$$
(3.11)

Applying Itô's formula to $e^t V(x_1, x_2)$ yields that

$$d[e^{t}V(x_{1}, x_{2})] = e^{t}V(x_{1}, x_{2}) dt + e^{t} dV(x_{1}, x_{2})$$

$$\leq \mathcal{L}e^{t} dt + p\sigma_{1}x_{1}^{p} dB_{1}(t) + p\sigma_{2}x_{2}^{p} dB_{2}(t).$$
(3.12)

Integrating from 0 to t and then taking the expectation of both sides yield that

$$E[V(x_1(t), x_2(t))] \le \exp(-t)[V(x_1(0), x_2(0)) + \mathcal{L}(\exp(t) - 1)].$$
(3.13)

This gives that

$$\limsup_{t \to \infty} E |x(t)|^p \le 2^{p-1} \limsup_{t \to \infty} E [V(x_1(t), x_2(t))] \le 2^{p-1} \mathcal{L} \triangleq K(p).$$
(3.14)

This completes the proof of Theorem 3.2.

Theorem 3.3 If $C_0^* < (r_{i0} - 0.5\sigma_i^2)/r_{i1}$, then system (1.4) is stochastically permanent.

Proof Taking arbitrary $0 < \varepsilon < 1$, we first claim that there is a constant $\alpha > 0$ such that

$$\liminf_{t \to \infty} P\{x_i(t) \ge \alpha\} \ge 1 - \varepsilon.$$
(3.15)

Since $C_0^* < (r_{i0} - 0.5\sigma_i^2)/r_{i1}$, we can choose a constant m > 0 such that

$$r_{i0} - 0.5\sigma_i^2 - r_{i1}C_0^* - 0.5m\sigma_i^2 > 0, \quad i = 1, 2.$$
(3.16)

Define

$$V_1(x_1, x_2) = m^{-1} \sum_{i=1}^{2} \left(1 + x_i^{-1} \right)^m, \quad (x_1, x_2) \in \mathbb{R}^2_+.$$
(3.17)

By Itô's formula we derive that

$$\begin{aligned} dV_1(x_1, x_2) \\ &= \left(1 + x_1^{-1}\right)^{m-1} \left[-x_1^{-2} dx_1 + x_1^{-3} (dx_1)^2\right] + 0.5(m-1) \left(1 + x_1^{-1}\right)^{m-2} \sigma_1^2 x_1^{-2} dt \\ &+ \left(1 + x_2^{-1}\right)^{m-1} \left[-x_2^{-2} dx_2 + x_2^{-3} (dx_2)^2\right] + 0.5(m-1) \left(1 + x_2^{-1}\right)^{m-2} \sigma_1^2 x_2^{-2} dt \\ &= \left(1 + x_1^{-1}\right)^{m-2} \left\{ \left(1 + x_1^{-1}\right) \left[-x_1^{-1} \left(r_{10} - r_{11}C_0(t) - a_1x_1 + \frac{c_2x_2}{1 + b_2x_2}\right) dt - \sigma_1x_1^{-1} dB_1(t) \right. \\ &+ \left. \sigma_1^2 x_1^{-1} dt \right] + 0.5(m-1) \sigma_1^2 x_1^{-2} dt \right\} + \left(1 + x_2^{-1}\right)^{m-2} \left\{ \left(1 + x_2^{-1}\right) \left[-x_2^{-1} \left(r_{20} - r_{21}C_0(t)\right) dt - \sigma_1x_1 + \frac{c_2x_2}{1 + b_2x_2}\right) dt - \sigma_1x_1 + \frac{c_2x_2}{1 + b_2x_2} dt \right\} \end{aligned}$$

$$-a_{2}x_{2} + \frac{c_{1}x_{1}}{1+b_{1}x_{1}} dt - \sigma_{2}x_{2}^{-1} dB_{2}(t) + \sigma_{2}^{2}x_{2}^{-1} dt + 0.5(m-1)\sigma_{2}^{2}x_{2}^{-2} dt$$

$$= \left(1 + x_{1}^{-1}\right)^{m-2} \left\{ -x_{1}^{-2} \left(r_{10} - r_{11}C_{0}(t) + \frac{c_{2}x_{2}}{1+b_{2}x_{2}} - 0.5\sigma_{1}^{2} - 0.5m\sigma_{1}^{2}\right) + x_{1}^{-1} \left(-r_{10}\right)$$

$$+ r_{11}C_{0}(t) - \frac{c_{2}x_{2}}{1+b_{2}x_{2}} + \sigma_{1}^{2} + a_{1} + a_{1} dt + \left(1 + x_{2}^{-1}\right)^{m-2} \left\{ -x_{2}^{-2} \left(r_{20} - r_{21}C_{0}(t)\right)$$

$$+ \frac{c_{1}x_{1}}{1+b_{1}x_{1}} - 0.5\sigma_{2}^{2} - 0.5m\sigma_{2}^{2} + x_{2}^{-1} \left(-r_{20} + r_{21}C_{0}(t) - \frac{c_{1}x_{1}}{1+b_{1}x_{1}} + \sigma_{2}^{2} + a_{2}\right)$$

$$+ a_{2} dt - \left(1 + x_{1}^{-1}\right)^{m-1}\sigma_{1}x_{1}^{-1} dB_{1}(t) - \left(1 + x_{2}^{-1}\right)^{m-1}\sigma_{2}x_{2}^{-1} dB_{2}(t)$$

$$\leq \left(1 + x_{1}^{-1}\right)^{m-2} \left\{ -x_{1}^{-2} \left(r_{10} - r_{11}C_{0}^{*} - \varepsilon - 0.5\sigma_{1}^{2} - 0.5m\sigma_{1}^{2}\right) + x_{1}^{-1} \left(-r_{10} + r_{11}C_{0}^{*} + \varepsilon + \sigma_{1}^{2} + a_{1}\right) dt + \left(1 + x_{2}^{-1}\right)^{m-2} \left\{ -x_{2}^{-2} \left(r_{20} - r_{21}C_{0}^{*} - \varepsilon - 0.5\sigma_{2}^{2} - 0.5m\sigma_{2}^{2}\right)$$

$$+ x_{2}^{-1} \left(-r_{20} + r_{21}C_{0}^{*} + \varepsilon + \sigma_{2}^{2} + a_{2}\right) + a_{2} dt - \left(1 + x_{1}^{-1}\right)^{m-1}\sigma_{1}x_{1}^{-1} dB_{1}(t)$$

$$- \left(1 + x_{2}^{-1}\right)^{m-1}\sigma_{2}x_{2}^{-1} dB_{2}(t).$$
(3.18)

Let k be sufficiently small to satisfy

$$0 < \frac{k}{m} < r_{i0} - r_{i1}C_0^* - \varepsilon - 0.5\sigma_i^2 - 0.5m\sigma_i^2, \quad i = 1, 2.$$
(3.19)

We define

$$V_2(t, x_1, x_2) = e^{kt} V_1(x_1, x_2).$$
(3.20)

By Itô's formula again,

$$\begin{aligned} dV_{2}(x_{1},x_{2}) \\ &= ke^{kt}V_{1}dt + e^{kt}dV_{1} \\ &\leq e^{kt}\left(1+x_{1}^{-1}\right)^{m-2}\left\{km^{-1}\left(1+x_{1}^{-1}\right)^{2} - x_{1}^{-2}\left(r_{10}-r_{11}C_{0}^{*}-\varepsilon-0.5\sigma_{1}^{2}-0.5m\sigma_{1}^{2}\right) \\ &+ x_{1}^{-1}\left(-r_{10}+r_{11}C_{0}^{*}+\varepsilon+\sigma_{1}^{2}+a_{1}\right)+a_{1}\right\}dt + e^{kt}\left(1+x_{2}^{-1}\right)^{m-2}\left\{km^{-1}\left(1+x_{2}^{-1}\right)^{2}\right. \\ &- x_{2}^{-2}\left(r_{20}-r_{21}C_{0}^{*}-\varepsilon-0.5\sigma_{2}^{2}-0.5m\sigma_{2}^{2}\right)+x_{2}^{-1}\left(-r_{20}+r_{21}C_{0}^{*}+\varepsilon+\sigma_{2}^{2}+a_{2}\right) \\ &+ a_{2}\right\}dt - \left(1+x_{1}^{-1}\right)^{m-1}\sigma_{1}x_{1}^{-1}e^{kt}dB_{1}(t) - \left(1+x_{2}^{-1}\right)^{m-1}\sigma_{2}x_{2}^{-1}e^{kt}dB_{2}(t) \end{aligned}$$

$$&= e^{kt}\left(1+x_{1}^{-1}\right)^{m-2}\left\{-x_{1}^{-2}\left(r_{10}-r_{11}C_{0}^{*}-\varepsilon-0.5\sigma_{1}^{2}-0.5m\sigma_{1}^{2}-\frac{k}{m}\right)+x_{1}^{-1}\left(-r_{10}\right) \\ &+ r_{11}C_{0}^{*}+\varepsilon+\sigma_{1}^{2}+a_{1}+\frac{2k}{m}\right)+a_{1}+\frac{k}{m}\right\}dt + e^{kt}\left(1+x_{2}^{-1}\right)^{m-2}\left\{-x_{2}^{-2}\left(r_{20}-r_{21}C_{0}^{*}\right) \\ &- \varepsilon-0.5\sigma_{2}^{2}-0.5m\sigma_{2}^{2}-\frac{k}{m}\right)+x_{2}^{-1}\left(-r_{20}+r_{21}C_{0}^{*}+\varepsilon+\sigma_{2}^{2}+a_{2}+\frac{2k}{m}\right)+a_{2}\right. \\ &+ \frac{k}{m}\right\}dt - \left(1+x_{1}^{-1}\right)^{m-1}\sigma_{1}x_{1}^{-1}e^{kt}dB_{1}(t) - \left(1+x_{2}^{-1}\right)^{m-1}\sigma_{2}x_{2}^{-1}e^{kt}dB_{2}(t) \\ &= e^{kt}J(x_{1},x_{2})dt - \left(1+x_{1}^{-1}\right)^{m-1}\sigma_{1}x_{1}^{-1}e^{kt}dB_{1}(t) - \left(1+x_{2}^{-1}\right)^{m-1}\sigma_{2}x_{2}^{-1}e^{kt}dB_{2}(t). \quad (3.21) \end{aligned}$$

Clearly, $J(x_1, x_2)$ is upper bounded for $(x_1, x_2) \in \mathbb{R}^2_+$, that is, $\mathcal{J} \triangleq \sup_{x \in \mathbb{R}^2_+} J(x_1, x_2) < \infty$. As a result, we have

$$dV_2(x_1, x_2) \le \mathcal{J}e^{kt} dt - \left(1 + x_1^{-1}\right)^{m-1} \sigma_1 x_1^{-1} e^{kt} dB_1(t) - \left(1 + x_2^{-1}\right)^{m-1} \sigma_2 x_2^{-1} e^{kt} dB_2(t).$$
(3.22)

Integrating both sides and then taking the expectations, we obtain that

$$E[V_2(x_1, x_2)] = e^{kt} E[V_1(x_1, x_2)] \le V_1[x_1(0), x_2(0)] + \frac{\mathcal{J}}{k} (e^{kt} - 1),$$
(3.23)

that is,

$$\limsup_{t \to \infty} E[V_1(x_1, x_2)] \le \frac{\mathcal{J}}{k}.$$
(3.24)

On the other hand,

$$\limsup_{t \to \infty} E[x_i^{-m}(t)] \leq \limsup_{t \to \infty} E\left[\left(1 + \frac{1}{x_i}\right)^m\right]$$
$$\leq m \limsup_{t \to \infty} E[V_1(x_1, x_2)]$$
$$\leq \frac{m\mathcal{J}}{k} \triangleq \delta.$$
(3.25)

For arbitrary $\varepsilon \in (0, 1)$, let $\alpha = (\frac{\varepsilon}{\delta})^{\frac{1}{m}}$. By Chebyshev's inequality we have

$$P\{x_i(t) < \alpha\} = P\{x_i^{-m}(t) > \alpha^{-m}\} \le \frac{E[x_i^{-m}(t)]}{\alpha^{-m}} = \alpha^m E[x_i^{-m}(t)],$$
(3.26)

which gives that

$$\limsup_{t \to \infty} P\{x_i(t) < \alpha\} \le \alpha^m \delta = \varepsilon,$$
(3.27)

that is,

$$\liminf_{t \to \infty} P\{x_i(t) \ge \alpha\} \ge 1 - \varepsilon.$$
(3.28)

In the following, we prove that for arbitrary $\varepsilon \in (0, 1)$, there is a constant $\beta > 0$ such that

$$\liminf_{t \to \infty} P\{x_i(t) \le \beta\} \ge 1 - \varepsilon.$$
(3.29)

Let $\beta = \left[\frac{K}{\varepsilon}\right]^{\frac{1}{p}}$. Then by Chebyshev's inequality and Theorem 3.2 we have

$$P\{x_i(t) > \beta\} = P\{x_i^p(t) > \beta^p\} \le \frac{E[x_i^p(t)]}{\beta^p},$$
(3.30)

which implies that

$$\limsup_{t \to \infty} P\{x_i(t) > \beta\} \le \frac{K}{\beta^p} = \varepsilon.$$
(3.31)

Consequently,

$$\liminf_{t \to \infty} P\{x_i(t) \le \beta\} \ge 1 - \varepsilon.$$
(3.32)

This completes the proof of Theorem 3.3.

Remark 3.1 Theorem 3.3, which directly measures the population size $x_i(t)$, indicates that the population size will neither too small nor too large with large probability when the time is sufficiently large.

Theorem 3.4 If $(C_0)_* > (r_{i0} - 0.5\sigma_i^2 + \frac{c_j}{b_j})/r_{i1}$, $i, j = 1, 2, i \neq j$, then population x_i goes to extinction with probability one.

Proof Applying Itô's formula yields that

$$d\ln x_{i} = \frac{1}{x_{i}} dx_{i} - \frac{1}{2x_{i}^{2}} (dx_{i})^{2}$$
$$= \left[r_{i0} - 0.5\sigma_{i}^{2} - r_{i1}C_{0}(t) - a_{i}x_{i} + \frac{c_{j}x_{j}}{1 + b_{j}x_{j}} \right] dt + \sigma_{i} dB_{i}(t).$$
(3.33)

Integrating both sides yields that

$$\frac{\ln x_i(t) - \ln x_i(0)}{t} = \left(r_{i0} - 0.5\sigma_i^2\right) - r_{i1}\langle C_0(t) \rangle - a_i \langle x_i(t) \rangle + \left(\frac{c_j x_j(t)}{1 + b_j x_j(t)}\right) + \sigma_i \frac{B_i(t)}{t}.$$
 (3.34)

Letting $t \to \infty$ and applying the strong law of large numbers, we obtain that

$$\limsup_{t \to \infty} \frac{\ln x_i(t)}{t} \le r_{i0} - 0.5\sigma_i^2 + \frac{c_j}{b_j} - r_{i1} \langle C_0 \rangle_* < 0.$$
(3.35)

In other words, $\lim_{t\to\infty} x_i(t) = 0$. This completes the proof of Theorem 3.4.

Remark 3.2 Theorem 3.4 indicates the worst case that the population will go to extinction almost surely.

Theorem 3.5 If $(C_0)_* < (r_{i0} - 0.5\sigma_i^2 + \frac{c_j}{b_j})/r_{i1}$, $i, j = 1, 2, i \neq j$, then population x_i is weakly persistent almost surely.

Proof We denote $S := \{\omega : \limsup_{t\to\infty} x_i(t,\omega) = 0\}$ and assume that P(S) > 0. Then for all $\omega \in S$, we have $\limsup_{t\to\infty} x_i(t,\omega) = 0$. For arbitrary small ε satisfying $0 < \varepsilon < 1$, there exists $T(\omega)$ such that

$$x_i(t,\omega) \le \varepsilon \quad \text{for } t \ge T(\omega).$$
 (3.36)

It then follows that

$$\limsup_{t \to \infty} \frac{\ln x_i(t) - \ln x_i(0)}{t} \le 0.$$
(3.37)

By the continuity of $x_i(t, \omega)$ there must be a constant \tilde{K} such that $x_i(t, \omega) \leq \tilde{K}$ for $0 \leq t \leq T(\omega)$. On the other hand,

$$\frac{1}{t} \int_{0}^{t} x_{i}(s,\omega) \, ds = \frac{1}{t} \int_{0}^{T(\omega)} x_{i}(s,\omega) \, ds + \frac{1}{t} \int_{T(\omega)}^{t} x_{i}(s,\omega) \, ds$$
$$\leq \tilde{K} \frac{T(\omega)}{t} + \varepsilon \frac{t - T(\omega)}{t} \tag{3.38}$$

for sufficient large *t*. Since ε is arbitrarily small, we obtain that

$$\limsup_{t \to \infty} \frac{1}{t} \int_0^t x_i(s,\omega) \, ds \le 0. \tag{3.39}$$

Since $x_i(t) > 0$, we have $\liminf_{t\to\infty} \frac{1}{t} \int_0^t x_i(s,\omega) \, ds \ge 0$, and thus

$$\lim_{t \to \infty} \frac{1}{t} \int_0^t x_i(s,\omega) \, ds = 0. \tag{3.40}$$

Substituting (3.40) into (3.34) and using the strong law of large numbers, we deduce the contradiction

$$\limsup_{t \to \infty} \frac{\ln x_i(t) - \ln x_i(0)}{t} \le r_i - 0.5\sigma_i^2 - r_{i1} \langle C_0 \rangle_* + \frac{c_j}{b_j} < 0.$$
(3.41)

This completes the proof of Theorem 3.5.

Remark 3.3 Theorem 3.5 admit the case that the population size is close to zero even if the time is sufficiently large. In this case the survival of species can be dangerous in reality. In addition, Theorems 3.4 and 3.5 reveal that $\langle C_0 \rangle_* = (r_{i0} - 0.5\sigma_i^2 + \frac{c_j}{b_j})/r_{i1}$ is the threshold between extinction and weak persistence.

4 The estimation of temporal averages

In this section, we further discuss the survival analysis of system (1.4) based on the temporal average of the population size $x_i(t)$.

Theorem 4.1 The solution of system (1.4) has the property that

$$\limsup_{t \to \infty} \frac{\ln x_i(t)}{\ln t} \le 1 \quad a.s.$$
(4.1)

Proof By Itô's formula,

$$d(e^{t} \ln x_{i}) = e^{t} \ln x_{i} dt + e^{t} d \ln x_{i}$$
$$= e^{t} \left[\ln x_{i} + r_{i0} - 0.5\sigma_{i}^{2} - r_{i1}C_{0}(t) - a_{i}x_{i} + \frac{c_{j}x_{j}}{1 + b_{j}x_{j}} \right] dt + \sigma_{i}e^{t} dB_{i}(t).$$
(4.2)

Integrating both sides yields that

$$e^{t}\ln x_{i}(t) - \ln x_{i}(0) = \int_{0}^{t} e^{s} \left[\ln x_{i} + r_{i0} - 0.5\sigma_{i}^{2} - r_{i1}C_{0}(s) - a_{i}x_{i} + \frac{c_{j}x_{j}}{1 + b_{j}x_{j}} \right] ds + N(t), \quad (4.3)$$

where $N(t) = \int_0^t \sigma_i e^s dB_i(s)$. Note that N(t) is a local martingale with quadratic variation

$$\langle N,N\rangle = \int_0^t \sigma_i^2 e^{2s} \, ds. \tag{4.4}$$

By the exponential martingale inequality we have

$$P\left\{\sup_{0\le t\le\gamma n} \left[N(t) - 0.5e^{-\gamma n} \langle N, N \rangle\right] > \theta e^{\gamma n} \ln n\right\} \le n^{-\theta},\tag{4.5}$$

where $\theta > 1$, and $\gamma > 0$ is arbitrary. By the Borel–Cantelli lemma, for almost all $\omega \in \Omega$, there is a random integer $n_0(\omega)$ such that for all $n \ge n_0(\omega)$ and $0 \le t \le \gamma n$,

$$N(t) \le 0.5e^{-\gamma n} \langle N, N \rangle + \theta e^{\gamma n} \ln n.$$
(4.6)

Substituting (4.6) into (4.3) yields that

$$e^{t} \ln x_{i}(t) - \ln x_{i}(0) \leq \int_{0}^{t} e^{s} \left[\ln x_{i} + r_{i0} - 0.5\sigma_{i}^{2} - r_{i1}C_{0}(s) - a_{i}x_{i} + \frac{c_{j}}{b_{j}} \right] ds$$

+ $0.5e^{-\gamma n} \int_{0}^{t} \sigma_{i}^{2}e^{2s} ds + \theta e^{\gamma n} \ln n$
= $\int_{0}^{t} e^{s} \left[\ln x_{i} + r_{i0} - 0.5\sigma_{i}^{2} - r_{i1}C_{0}(s) - a_{i}x_{i} + \frac{c_{j}}{b_{j}} + 0.5\sigma_{i}^{2}e^{s-\gamma n} \right] ds + \theta e^{\gamma n} \ln n.$ (4.7)

Since $0 \le t \le \gamma n$ and $x_i(t) > 0$, there is a constant ρ such that

$$\ln x_i + r_{i0} - 0.5\sigma_i^2 - r_{i1}C_0(t) - a_i x_i + \frac{c_j}{b_j} + 0.5\sigma_i^2 e^{t - \gamma n} \le \varrho,$$
(4.8)

that is, for $0 \le t \le \gamma n$, we have

$$e^{t}\ln x_{i}(t) - \ln x_{i}(0) \leq \varrho\left(e^{t} - 1\right) + \theta e^{\gamma n}\ln n.$$

$$(4.9)$$

Therefore, if $\gamma(n-1) \le t \le \gamma n$ and $n \ge n_0(\omega)$, then we obtain that

$$\frac{\ln x_i(t)}{\ln t} \le \frac{e^{-t} \ln x_i(0)}{\ln t} + \frac{\varrho(1 - e^{-t})}{\ln t} + \frac{\theta e^{-\gamma(n-1)} e^{\gamma n} \ln n}{\ln t},$$
(4.10)

which implies that

$$\limsup_{t \to \infty} \frac{\ln x_i(t)}{\ln t} \le \theta e^{\gamma}.$$
(4.11)

Then letting $\theta \to 1$ and $\gamma \to 0$ gives the required assertion (4.1). This completes the proof of Theorem 4.1.

Theorem 4.2 Let x(t) be a solution of system (1.4). If $r_{i0} - 0.5\sigma_i^2 - r_{i1}\langle C_0 \rangle^* > 0$, then the component $x_i(t)$ satisfies

$$\lim_{t \to \infty} \frac{\ln x_i(t)}{t} = 0 \quad a.s.$$
(4.12)

Proof We first show that $\limsup_{t\to\infty} \frac{\ln x_i(t)}{t} \leq 0$. From Theorem 4.1 we have that

$$\limsup_{t \to \infty} \frac{\ln x_i(t)}{t} = \limsup_{t \to \infty} \frac{\ln x_i(t)}{\ln t} \limsup_{t \to \infty} \frac{\ln t}{t}$$
$$\leq \limsup_{t \to \infty} \frac{\ln t}{t}$$
$$= 0. \tag{4.13}$$

In the following, we show that $\liminf_{t\to\infty} \frac{\ln x_i(t)}{t} \ge 0$. Since $\lim_{t\to\infty} \frac{\int_0^t \sigma_i \, dB_i(s)}{t} = 0$ a.s., for any $\varepsilon \in (0, 1)$, there exists a positive constant T such that

$$\left|\int_{s}^{t} \sigma_{i} dB_{i}(\tau)\right| \leq \left|\int_{0}^{t} \sigma_{i} dB_{i}(\tau)\right| + \left|\int_{0}^{s} \sigma_{i} dB_{i}(\tau)\right| \leq \varepsilon(t+s) \quad \text{a.s.}$$

$$(4.14)$$

for $t > s \ge T$. From equations (3.38) and (4.14) we have

$$\begin{aligned} x_{i}^{-1}(t) &\leq x_{i}^{-1}(T) \exp\left\{-\left(r_{i0} - 0.5\sigma_{i}^{2}\right)(t - T) + r_{i1}\int_{T}^{t}C_{0}(s)\,ds - \int_{T}^{t}\sigma_{i}\,dB_{i}(\tau)\right\} \\ &+ a_{i}\int_{T}^{t}\exp\left\{\left(r_{i0} - 0.5\sigma_{i}^{2}\right)(s - t) + r_{i1}\int_{s}^{t}C_{0}(\tau)\,d\tau - \int_{s}^{t}\sigma_{i}\,dB_{i}(\tau)\right\}\,ds \\ &\leq x_{i}^{-1}(T)\exp\left\{-\left(r_{i0} - 0.5\sigma_{i}^{2} - r_{i1}\langle C_{0}\rangle^{*} - r_{i1}\varepsilon\right)(t - T) + \varepsilon(t + T)\right\} \\ &+ a_{i}\int_{T}^{t}\exp\left\{\left(r_{i0} - 0.5\sigma_{i}^{2} - r_{i1}\langle C_{0}\rangle^{*} - r_{i1}\varepsilon\right)(s - t) + \varepsilon(t + s)\right\}\,ds. \end{aligned}$$
(4.15)

This gives that

$$e^{-2\varepsilon(t+T)}x_i^{-1}(t) \le x_i^{-1}(T)\exp\left\{-\left(r_{i0}-0.5\sigma_i^2-r_{i1}\langle C_0\rangle^*-r_{i1}\varepsilon\right)(t-T)-\varepsilon(t+T)\right\}$$
$$+a_i\int_T^t\exp\left\{\left(r_{i0}-0.5\sigma_i^2-r_{i1}\langle C_0\rangle^*-r_{i1}\varepsilon\right)(s-t)-\varepsilon(t-s)-2\varepsilon T\right\}ds$$
$$\le \bar{K},$$
(4.16)

that is, $x_i^{-1}(t) \leq \bar{K} e^{2\varepsilon(t+T)}$ almost surely. Therefore we obtain that

$$\frac{\ln x_i^{-1}(t)}{t} \le \frac{1}{t} \left[\ln \bar{K} + 2\varepsilon(t+T) \right] \quad \text{a.s.,}$$

$$(4.17)$$

which yields that

$$\limsup_{t \to \infty} \frac{\ln x_i^{-1}(t)}{t} \le 2\varepsilon \quad \text{a.s.}$$
(4.18)

In other words,

$$\liminf_{t \to \infty} \frac{\ln x_i(t)}{t} \ge -2\varepsilon \quad \text{a.s.}$$
(4.19)

Then from the arbitrariness of ε it follows that

$$\liminf_{t \to \infty} \frac{\ln x_i(t)}{t} \ge 0 \quad \text{a.s.}$$
(4.20)

This completes the proof of Theorem 4.2.

Theorem 4.3 Let x(t) be a solution of system (1.4). If $r_{i0} - 0.5\sigma_i^2 - r_{i1}\langle C_0 \rangle^* > 0$, then the component $x_i(t)$ has the following property:

$$\limsup_{t \to \infty} \frac{\int_0^t x_i(s) \, ds}{t} \le \frac{r_{i0} - 0.5\sigma_i^2 + \frac{c_j}{b_j} - r_{i1} \langle C_0 \rangle_*}{a_i} \quad a.s.$$
(4.21)

Moreover,

$$\liminf_{t \to \infty} \frac{\int_0^t x_i(s) \, ds}{t} \ge \frac{r_{i0} - 0.5\sigma_i^2 - r_{i1} \langle C_0 \rangle^*}{a_i}, \quad i = 1, 2, \ a.s.$$
(4.22)

that is, population x_i is strongly persistent in the mean almost surely.

Proof Recalling (3.34), we obtain that

$$\frac{\ln x_i(t) - \ln x_i(0)}{t} + a_i \langle x_i(t) \rangle \le \left(r_{i0} - 0.5\sigma_i^2 \right) - r_{i1} \langle C_0(t) \rangle + \frac{c_j}{b_j} + \sigma_i \frac{B_i(t)}{t}.$$
(4.23)

Then from Theorem 4.2 it follows that

$$\langle x_i(t) \rangle^* \le \frac{r_{i0} - 0.5\sigma_i^2 + \frac{c_j}{b_j} - r_{i1} \langle C_0 \rangle_*}{a_i}.$$
 (4.24)

On the other hand,

$$\frac{\ln x_i(t) - \ln x_i(0)}{t} + a_i \langle x_i(t) \rangle \ge \left(r_{i0} - 0.5\sigma_i^2 \right) - r_{i1} \langle C_0 \rangle + \sigma_i \frac{B_i(t)}{t}.$$
(4.25)

Similarly, we obtain that

$$\langle x_i(t) \rangle_* \ge \frac{r_{i0} - 0.5\sigma_i^2 - r_{i1} \langle C_0 \rangle^*}{a_i}.$$
 (4.26)

This completes the proof of Theorem 4.3.

Theorem 4.4 If $\langle C_0 \rangle_* < (r_{i0} - 0.5\sigma_i^2)/r_{i1}$, then the component $x_i(t)$ satisfies

$$\limsup_{t \to \infty} \langle x_i(t) \rangle > 0. \tag{4.27}$$

In other words, population x_i is weakly persistent in the mean almost surely.

Proof Recalling (4.25), we obtain that

$$\limsup_{t \to \infty} \frac{\ln x_i(t) - \ln x_i(0)}{t} + a_i \langle x_i(t) \rangle^* \ge (r_{i0} - 0.5\sigma_i^2) - r_{i1} \langle C_0 \rangle_*.$$
(4.28)

For all $\omega \in \{\omega : \limsup_{t\to\infty} \langle x_i(t,\omega) \rangle = 0\}$, we have $\langle x_i(t,\omega) \rangle^* = 0$. Then from Theorem 4.2 and (4.28) it follows that

$$0 \ge (r_{i0} - 0.5\sigma_i^2) - r_{i1} \langle C_0 \rangle_*, \tag{4.29}$$

which contradicts the assumption of Theorem 4.4. So we must have $\limsup_{t\to\infty} \langle x_i(t) \rangle > 0$. This completes the proof of Theorem 4.4.

5 Numerical simulations

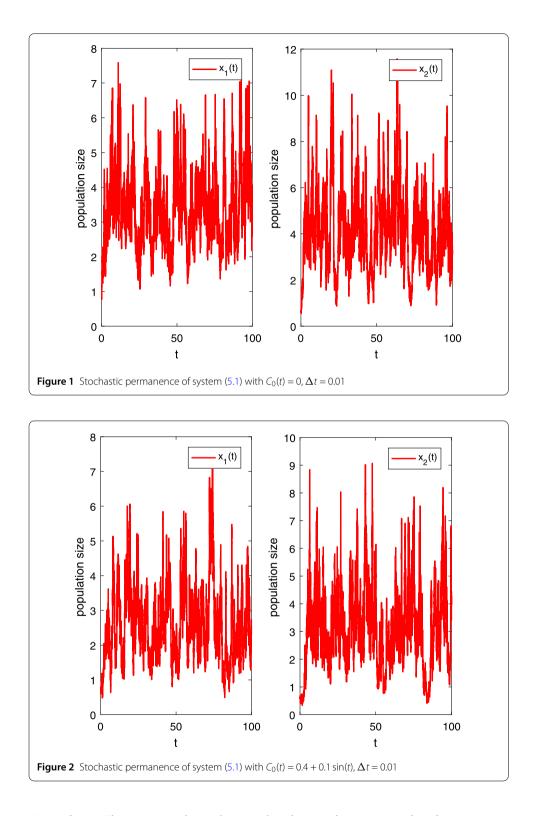
In this section, we present several specific examples to justify our theoretical results based on the Milstein method, which is mentioned by Higham [36]. For system (1.4), we assign $r_{10} = 1.6$, $r_{11} = 0.8$, $a_1 = 0.5$, $b_2 = 0.4$, $c_2 = 0.2$, $\sigma_1 = 0.6$, $r_{20} = 1.4$, $r_{21} = 0.7$, $a_2 = 0.4$, $b_1 = 0.5$, $c_1 = 0.4$, $\sigma_2 = 0.7$, and the initial value ($x_1(0), x_2(0)$) = (0.8, 0.6) and then consider the following discrete version:

$$\begin{cases} x_{1}(k+1) = x_{1}(k) + x_{1}(k)[r_{10} - r_{11}C_{0}(k\Delta t) - a_{1}x_{1}(k) + \frac{c_{2}x_{2}(k)}{1 + b_{2}x_{2}(k)}]\Delta t \\ + \sigma_{1}x_{1}(k)\sqrt{\Delta t}\xi_{1}(k) + 0.5\sigma_{1}^{2}x_{1}(k)[\xi_{1}^{2}(k) - 1]\Delta t, \\ x_{2}(k+1) = x_{2}(k) + x_{2}(k)[r_{20} - r_{21}C_{0}(k\Delta t) - a_{2}x_{2}(k) + \frac{c_{1}x_{1}(k)}{1 + b_{1}x_{1}(k)}]\Delta t \\ + \sigma_{2}x_{2}(k)\sqrt{\Delta t}\xi_{2}(k) + 0.5\sigma_{2}^{2}x_{2}(k)[\xi_{2}^{2}(k) - 1]\Delta t, \end{cases}$$
(5.1)

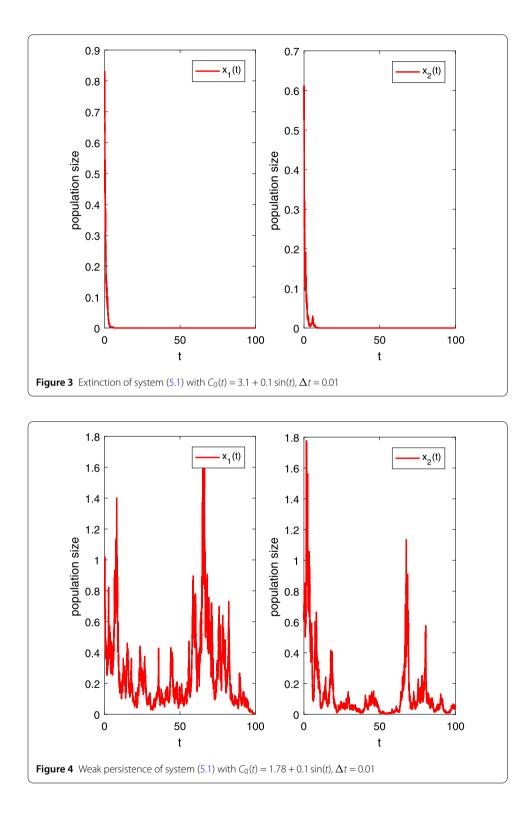
where $\xi_1(k)$ and $\xi_2(k)$ are independent standard Gaussian random variables. To reveal the effects of toxicant, we need to choose $C_0(t) = 0$ for comparison with system (5.1). By a simple calculation, $r_{10} = 1.6 > 0.5\sigma_1^2 = 0.18$, $r_{20} = 1.4 > 0.5\sigma_2^2 = 0.245$. Then each species can live alone in a nonpolluted environment; see Fig. 1. Theorem 3.1 shows that for any positive initial value, the solution of system (5.1) remains in R_+^2 almost surely. The nice property provides us a great opportunity to discuss the survival analysis of system (5.1), and so it is the fundamental conclusion in this paper.

Example 5.1 Theorem 3.3 shows that a tiny amount of toxicant cannot disrupt the stochastic permanence; in other words, the population size will be neither too small nor too large with large probability if the time is sufficiently large. Let $C_0(t) = 0.4 + 0.1 \sin(t)$, $\Delta t = 0.01$; by a simple calculation, $C_0^* = 0.5$, $(r_{10} - 0.5\sigma_1^2)/r_{11} = 1.775$, $(r_{20} - 0.5\sigma_2^2)/r_{21} = 1.65$, that is, $C_0^* < (r_{i0} - 0.5\sigma_i^2)/r_{i1}$. So population x_i is stochastically permanent by Theorem 3.3; see Fig. 2.

Example 5.2 Theorem 3.4 shows that a large amount of toxicant will force the species incline to extinction with probability one. Let $C_0(t) = 3.1 + 0.1 \sin(t)$, $\Delta t = 0.01$. It is easy to compute that $\langle C_0 \rangle_* = 3.1$, $(r_{10} - 0.5\sigma_1^2 + \frac{c_1}{b_1})/r_{11} = 2.4$, $(r_{20} - 0.5\sigma_2^2 + \frac{c_2}{b_2})/r_{21} = 2.79$, that is, $\langle C_0 \rangle_* > (r_{i0} - 0.5\sigma_i^2 + \frac{c_j}{b_j})/r_{i1}$. So population x_i goes to extinction almost surely; see Fig. 3.



Example 5.3 Theorem 3.5 admits the case that the population size is closed to zero even if the time is sufficiently large. Let $C_0(t) = 1.78 + 0.1 \sin(t)$, $\Delta t = 0.01$. Then by a straightforward calculation we have $\langle C_0 \rangle_* = 1.78$ and $\langle C_0 \rangle_* < (r_{i0} - 0.5\sigma_i^2 + \frac{c_j}{b_j})/r_{i1}$. By Theorem 3.5 population x_i is weakly persistent almost surely; see Fig. 4.



6 Discussion

In this paper, considering the fact of polluted environment, we propose and study a stochastic two-species cooperation model with functional response. However, we only consider the case that a coupling between species and toxicant is a linear function of concentration of toxicant present in the organism, that is, $r_{i0} - r_{i1}H_i(C_0) = r_{i0} - r_{i1}C_0(t)$, where

 $H_i(C_0)$ is the dose-response function of species x_i to the toxin. However, someone suggested that $H_i(C_0)$ should be nonlinear in many cases, such as a sigmoid dose-response curve (see [37, 38]). Liu and Wang [24] also introduced a more general case: $H_i(C_0)$ is a nondecreasing continuous function of C_0 with $H_i(0) = 0$. Clearly, our assumption is a particular case of this generalized condition. Moreover, the exogenous toxin cannot be continuous but rather of the impulse form, the growth rate of population is also inevitably affected by other environment noise such as Lévy jumps, and the regime-switching is other common random perturbation in the environment (see [39–41]). All these questions associated with the polluted environment are interesting topics to deserve further investigation, and we leave them for our future works.

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Availability of data and materials

Data sharing is not applicable to this paper as no datasets were generated or analyzed during the current study.

Competing interests

The author declares that he has no competing interests.

Author's contributions

The single author is responsible for the complete manuscript. The author read and approved the final manuscript.

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